Multiple Regression

Summary

The **Multiple Regression** procedure is designed to construct a statistical model describing the impact of a two or more quantitative factors X on a dependent variable Y. The procedure includes an option to perform a *stepwise regression*, in which a subset of the X variables is selected. The fitted model may be used to make predictions, including confidence limits and/or prediction limits. Residuals may also be plotted and influential observations identified.

The procedure contains additional options for transforming the data using either a Box-Cox or Cochrane-Orcutt transformation. The first option is useful for stabilizing the variability of the data, while the second is useful for handling time series data in which the residuals exhibit serial correlation.

Sample StatFolio: *multiple reg.sgp*

Sample Data:

The file *93cars.sgd* contains information on 26 variables for n = 93 makes and models of automobiles, taken from Lock (1993). The table below shows a partial list of 4 columns from that file:

Make	Model	MPG	Weight	Horsepower	Wheelbase	Drivetrain
		Highway				
Acura	Integra	31	2705	140	102	front
Acura	Legend	25	3560	200	115	front
Audi	90	26	3375	172	102	front
Audi	100	26	3405	172	106	front
BMW	535i	30	3640	208	109	rear
Buick	Century	31	2880	110	105	front
Buick	LeSabre	28	3470	170	111	front
Buick	Roadmaster	25	4105	180	116	rear
Buick	Riviera	27	3495	170	108	front
Cadillac	DeVille	25	3620	200	114	front
Cadillac	Seville	25	3935	295	111	front
Chevrolet	Cavalier	36	2490	110	101	front

A model is desired that can predict *MPG Highway* from *Weight*, *Horsepower*, *Wheelbase*, and *Drivetrain*.

Data Input

The data input dialog box requests the names of the columns containing the dependent variable Y and the independent variables X:

Multiple Regression	\mathbf{X}
Air Bags Cylinders Domestic Drive Train Engine Size Fueltank Horsepower Length Luggage Make Manual Max Price Mid Price Mid Price Mid Price Mid Price Min Price Model MPG City MPG Highway Passengers Rear seat Revs per Mile RPM Type	Dependent Variable: Independent Variables: Weight Weight^2 Horsepower Wheelbase Drive Train = "front" (Select:)
U Turn Space Weight 🛛 🖌	(Weights:)
Sort column names	
OK Cancel	Delete Transform Help

- **Y**: numeric column containing the *n* observations for the dependent variable Y.
- X: numeric columns containing the *n* values for the independent variables X. Either column names or STATGRAPHICS expressions may be entered.
- Select: subset selection.
- Weight: an optional numeric column containing weights to be applied to the squared residuals when performing a weighted least squares fit.

In the example, note the use of the expression *Weight*² to add a second-order term involving the weight of the vehicle. This was added after examining an X-Y plot that showed significant curvature with respect to *Weight*. The categorical factor *Drivetrain* has also be introduced into the model through the Boolean expression *Drivetrain="front"*, which sets up an indicator variable that takes the value 1 if true and 0 if false. The model to be fit thus takes the form:

$$MPG Highway = \beta_0 + \beta_1 Weight + \beta_2 Weight^2 + \beta_3 Horsepower + \beta_4 Wheelbase + \beta_5 X_5$$
(1)

where

 $X_{5} = \begin{cases} 1 & \text{if } Drivetrain = front \\ 0 & Drivetrain = rear \end{cases}$

(2)

Analysis Summary

The Analysis Summary shows information about the fitted model.

Multiple R	ogrocs	ion - MP(1 Hi	ahway		
Dependent vari			, 111	<u>guway</u>		
		1		Standard	T	
Parameter		Estimate		Error	Statistic	P-Value
CONSTANT		49.8458		10.5262	4.73539	0.0000
Weight		-0.0273685		0.00530942	-5.1547	0.0000
Weight^2		0.00000261	405	8.383E-7	3.11827	0.0025
Horsepower		0.0145764		0.009668	1.50769	0.1353
Wheelbase		0.338687		0.103479	3.273	0.0015
Drive Train="f	front"	0.632343		0.73879	0.855918	0.3944
Analysis of Va	riance					
Source	Sum of	Squares	Df	Mean Square	F-Ratio	P-Value
Model	1902.1	8	5	380.435	46.41	0.0000
Residual	713.13	6	87	8.19696		
Total (Corr.)	2615.3	1	92			
R-squared = 72 R-squared (adju Standard Error Mean absolute Durbin-Watson Lag 1 residual a	usted for of Est. = error = 2 statistic	d.f.) = 71.16 2.86303 2.13575 = 1.685 (P=).060	1)		

Included in the output are:

• Variables: identification of the dependent variable. The general form of the model is

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k$$
(3)

where *k* is the number of independent variables.

• **Coefficients:** the estimated coefficients, standard errors, t-statistics, and P values. The estimates of the model coefficients can be used to write the fitted equation, which in the example is

$$MPG Highway = 49.8458 - 0.0273685*Weight + 0.00000261405*Weight^{2} + 0.0145764*Horsepower + 0.338687*Wheelbase + 0.632343*Drive Train="front"$$
(4)

The t-statistic tests the null hypothesis that the corresponding model parameter equals 0, based on the Type 3 sums of squares (the extra sums of squares attributable to each variable if it is entered into the model last). Large P-Values (greater than or equal to 0.05 if operating at the 5% significance level) indicate that a term can be dropped without significantly degrading the model *provided all of the other variables remain in the model*. In the current © 2009 by StatPoint Technologies, Inc. Multiple Regression - 3

case, both *Horsepower* and *Drivetrain* are not significant. Thus, either variable (but not necessarily both) could be dropped from the model without hurting its predictive power significantly.

- Analysis of Variance: decomposition of the variability of the dependent variable Y into a model sum of squares and a residual or error sum of squares. Of particular interest is the F-test and its associated P-value, which tests the statistical significance of the fitted model. A small P-Value (less than 0.05 if operating at the 5% significance level) indicates that a significant relationship of the form specified exists between Y and the independent variables. In the sample data, the model is highly significant.
- Statistics: summary statistics for the fitted model, including:

R-squared - represents the percentage of the variability in Y which has been explained by the fitted regression model, ranging from 0% to 100%. For the sample data, the regression has accounted for about 72.7% of the variability in the miles per gallon. The remaining 27.3% is attributable to deviations from the model, which may be due to other factors, to measurement error, or to a failure of the current model to fit the data adequately.

Adjusted R-Squared – the R-squared statistic, adjusted for the number of coefficients in the model. This value is often used to compare models with different numbers of coefficients.

Standard Error of Est. – the estimated standard deviation of the residuals (the deviations around the model). This value is used to create prediction limits for new observations.

Mean Absolute Error – the average absolute value of the residuals.

Durbin-Watson Statistic – a measure of serial correlation in the residuals. If the residuals vary randomly, this value should be close to 2. A small P-value indicates a non-random pattern in the residuals. For data recorded over time, a small P-value could indicate that some trend over time has not been accounted for. In the current example, the P-value is greater than 0.05, so there is not a significant correlation at the 5% significance level.

Lag 1 Residual Autocorrelation – the estimated correlation between consecutive residuals, on a scale of -1 to 1. Values far from 0 indicate that significant structure remains unaccounted for by the model.

Multiple Regression Opt	tions				X
Fitting Procedure C Ordinary Least Squares Forward Stepwise Select Backward Stepwise Select Box-Cox Optimization C Cochrane-Orcutt Optimi	ction lection	Transfor Power: 1.0 Addend 0 Autocor 0			OK Cancel Help
Constant in Model	Stepwise Ro Selection F-Ratio P-Valu Max. Steps:	Criterion –	F-to-Enter: 4.0 P-to-Enter: 0.05 all steps	F-to-Ren 4.0 P-to-Ren 0.05	

Analysis Options

- Fitting Procedure specifies the method used to fit the regression model. The options are:
 - Ordinary Least Squares fits a model using all of the independent variables.
 - Forward Stepwise Selection performs a forward stepwise regression. Beginning with a model that includes only a constant, the procedure brings in variables one at a time provided that they will be statistically significant once added. Variables may also be removed at later steps if they are no longer statistically significant.
 - Backward Stepwise Selection performs a backward stepwise regression. Beginning with a model that includes all variables, the procedure removes variables one at a time if they are not statistically significant. Removed variables may also be added to the model at later steps if they become statistically significant.
 - Box-Cox Optimization fits a model involving all of the independent variables. The dependent variable, however, is modified by raising it to a power. The method of Box and Cox is used to determine the optimum power. Box-Cox transformations are a way of dealing with situations in which the deviations from the regression model do not have a constant variance.
 - Cochrane-Orcutt Optimization fits a model involving all of the independent variables. However, the least squares procedure is modified to allow for autocorrelation between successive residuals. The value of the lag 1 autocorrelation coefficient is determined using the method of Cochrane and Orcutt. The Cochrane-Orcutt transformation is a method for dealing with situations in which the model residuals are not independent.

- **Constant in model** If this option is not checked, the constant term β_0 will be omitted from the model. Removing the constant term allows for regression through the origin.
- **Power** Specifies the power to which the dependent variable is raised. The default value of 1.0 implies no power transformation.
- Addend Specifies an amount that is added to the dependent variable before it is raised to the specified power.
- Autocorrelation Specifies the lag 1 autocorrelation of the residuals. The default value of 0.0 implies that the residuals are assumed to be independent. If the Cochrane-Orcutt procedure is used, this value provides the starting value for the procedure.
- Selection Criterion If performing a forward or backward stepwise regression, this specifies whether variable entry and removal should be based on the F-ratio or its associated P-value.

=

- **F-to-Enter** In a stepwise regression, variables will be entered into the model at a given step if their F values are greater than or equal to the *F-to-Enter* value specified.
- **F-to-Remove** In a stepwise regression, variables will be removed from the model at a given step if their F values are less than the *F-to-Remove* value specified.
- **P-to-Enter** In a stepwise regression, variables will be entered into the model at a given step if their P values are less than or equal to the *P-to-Enter* value specified.
- **P-to-Remove** In a stepwise regression, variables will be removed from the model at a given step if their P values are greater than the *P-to-Remove* value specified.
- Max Steps maximum number of steps permitted when doing a stepwise regression.
- **Display** whether to display the results at each step when doing a stepwise regression.

Example – Stepwise Regression

The model fit to the automobile data showed 2 insignificant variables. To remove them from the model, *Analysis Options* may be used to perform either a forward stepwise selection or a backward stepwise selection.

• Forward selection – Begins with a model involving only a constant term and enters one variable at a time based on its statistical significance if added to the current model. At each step, the algorithm brings into the model the variable that will be the most statistically significant if entered. Selection of variables is based on either an *F*-to-enter test or a *P*-to-enter test. In the former case, as long as the most significant variable has an F value greater or equal to that specified on the *Analysis Summary* dialog box, it will be brought into the model. When no variable has a large enough F value, variable selection stops. In addition, variables brought into the model early in

the procedure may be removed later if their F value falls below the *F-to-remove* criterion.

• **Backward selection** – Begins with a model involving all the variables specified on the data input dialog box and removes one variable at a time based on its statistical significance in the current model. At each step, the algorithm removes from the model the variable that is the least statistically significant. Removal of variables is based on either an *F-to-remove* test or a *P-to-enter* test. In the former case, if the least significant variable has an F value less than that specified on the *Analysis Summary* dialog box, it will be removed from the model. When all remaining variables have large F values, the procedure stops. In addition, variables removed from the model early in the procedure may be re-entered later if their F values reach the *F-to-enter* criterion.

In the current example, a backwards selection procedure yields the following:

```
Stepwise regression
Method: backward selection
F-to-enter: 4.0
F-to-remove: 4.0
Step 0:
5 variables in the model. 87 d.f. for error.
R-squared = 72.73% Adjusted R-squared = 71.17% MSE = 8.19696
Step 1:
Removing variable Drive Train="front" with F-to-remove =0.732595
4 variables in the model. 88 d.f. for error.
R-squared = 72.50% Adjusted R-squared = 71.25% MSE = 8.17206
Step 2:
Removing variable Horsepower with F-to-remove =2.22011
3 variables in the model. 89 d.f. for error.
R-squared = 71.81% Adjusted R-squared = 70.86% MSE = 8.28409
Final model selected.
```

In the first step, *Drivetrain* is removed since it is the least significant. At the second step, *Horsepower* is removed. The algorithm then stops, since all remaining variables have F-to-remove values greater than 4, and all previously removed variables have F-to-enter values less than 4.

The reduced model is summarized below:

Multiple R	egression - MF	PG Highway		
Dependent vari	able: MPG Highwa	у		
Independent va	riables:			
Weight				
Weight^2				
Horsepower				
Wheelbase				
Drive Train-	="front"			
r	1			-
		Standard	Т	
Parameter	Estimate	Error	Statistic	P-Value
CONSTANT	51.8628	10.2179	5.07569	0.0000
Weight	-0.0245435	0.00506191	-4.84867	0.0000
Weight^2	0.00000236841	8.25606E-7	2.86869	0.0051
Wheelbase	0.28345	0.0899993	3.14947	0.0022

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Source	Sum of Squares	Df	Mean Square	F-Ratio	P-Value
Model	1878.03	3	626.009	75.57	0.0000
Residual	737.284	89	8.28409		
Total (Corr.)	2615.31	92			

R-squared = 71.809 percent R-squared (adjusted for d.f.) = 70.8587 percent Standard Error of Est. = 2.87821Mean absolute error = 2.19976Durbin-Watson statistic = 1.67296 (P=0.0558) Lag 1 residual autocorrelation = 0.162386

NOTE: from here forward in this document, the results will be based on the reduced model without *Drivetrain* or *Wheelbase*.

Example – Box-Cox Transformation

If it is suspected that the variability of Y changes as its level changes, it is useful to consider performing a transformation on Y. The Box-Cox transformations are of the general form

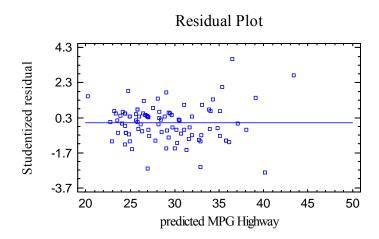
$$Y' = \left(Y + \lambda_2\right)^{\lambda_1} \tag{5}$$

in which the data is raised to a power λ_1 after shifting it a certain amount λ_2 . Often, the shift parameter λ_2 is set equal to 0. This class includes square roots, logarithms, reciprocals, and other common transformations, depending on the power. Examples include:

Power	Transformation	Description
$\lambda_1 = 2$	$Y' = Y^2$	square
$\lambda_1 = 1$	Y' = Y	untransformed data
$\lambda_1 = 0.5$	$Y' = \sqrt{Y}$	square root
$\lambda_1 = 0.333$	$Y' = \sqrt[3]{Y}$	cube root
$\lambda_1 = 0$	$Y' = \ln(Y)$	logarithm
$\lambda_1 = -0.5$	$Y' = \frac{1}{\sqrt{Y}}$	inverse square root
$\lambda_1 = -1$	$Y' = \frac{1}{Y}$	reciprocal

Using *Analysis Options*, you can specify the values for λ_1 or λ_2 , or specify just λ_2 and have the program find an optimal value for λ_1 using the methods proposed by Box and Cox (1964).

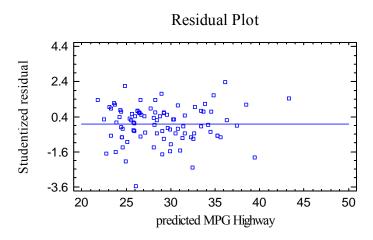
For the sample data, a plot of the residuals versus predicted values does show some change in variability as the predicted value changes:



The smaller cars tend to be somewhat more variable than the larger cars. Asking the program to optimize the Box-Cox transformation yields:

Multiple R	egression -	MPC	; Hi	ghv	vay			
Dependent vari	able: MPG Hig	hway						
Independent va	riables:	-						
Weight								
Weight^2								
Wheelbase								
Box-Cox transf	ormation applie	-		-0.4		= 0.0		1
			dard		Т			l
Parameter	Estimate	Erro			Statistic	P-Value		
CONSTANT	230.703	9.37	335		24.6126	0.0000		l
Weight	-0.0129299	0.00	4643:	53	-2.78451	0.0065		
Weight^2	6.18885E-7	7.57	367E	-7	0.817153	0.4160		
Wheelbase	0.229684	0.08	2560	6	2.782	0.0066		
Analysis of Va	riance							
Source	Sum of Square	es	Df	Me	ean Square	F-Ratio	P-Value	?
Model	1568.28		3	522	2.761	74.99	0.0000	
Residual	620.444		89	6.9	7128			
Total (Corr.)	2188.73		92					
· · · · · ·	•							
R-squared $= 71$.6528 percent							
R-squared (adju	usted for $d.f.$ =	70.69	72 pe	rcent	t			
Standard Error	of Est. $= 2.6403$	32						
Mean absolute	error = 2.08197							
Durbin-Watson	statistic $= 1.70$	034 (F	P=0.0	727)				
Lag 1 residual a	autocorrelation	= 0.14	8826					

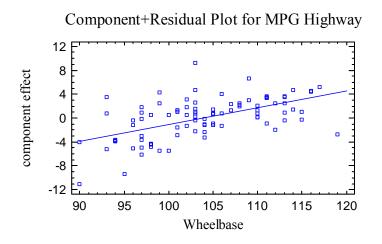
Apparently, an inverse square root of *MPG Highway* improves the properties of the residuals, as illustrated in the new residual plot:



Note: some caution is necessary here, however, since the transformation may be heavily influenced by one or two outliers. To simplify the discussion that follows, however, the rest of this document will work with the untransformed model.

Component Effects Plot

Plotting a multiple regression model is not as easy as plotting a simple regression model, since the space of the X variables is multi-dimensional. One useful way to illustrate the results is through the *Component Effects Plot*, which plots of the portion of the fitted regression model corresponding to any single variable.



The line on the plot is defined by

$$\hat{\boldsymbol{\beta}}_{j}\left(\boldsymbol{x}_{j}-\overline{\boldsymbol{x}}_{j}\right) \tag{6}$$

where $\hat{\beta}_{j}$ is the estimated regression coefficient for variable *j*, *x_j* represents the value of variable *j* as plotted on the horizontal axis, and \bar{x}_{j} is the average value of the selected independent variable amongst the *n* observations used to fit the model. You can judge the importance of a factor by noting how much the component effect changes over the range of the selected variable.

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For example, as *Wheelbase* changes from 90 to 120, the component effect changes from about -4 to +4. This implies that differences in *Wheelbase* account for a swing of about 8 miles per gallon.

The points on the above plot represent each of the n = 93 automobiles in the dataset. The vertical positions are equal to the component effect plus the residual from the fitted model. This allows you to gage the relative importance of a factor compared to the residuals. In the above plot, some of the residuals are as large if not larger than the effect of *Wheelbase*, indicating that other important factors may be missing from the model.

Pane Options

Component Effects Plot Options	<
Plot versus: Weight Weight^2 Horsepower Wheelbase Drive Train = "front"	
OK Cancel Help	

• Plot versus: the factor used to define the component effect.

Conditional Sums of Squares

The *Conditional Sums of Squares* pane displays a table showing the statistical significance of each coefficient in the model as it added to the fit:

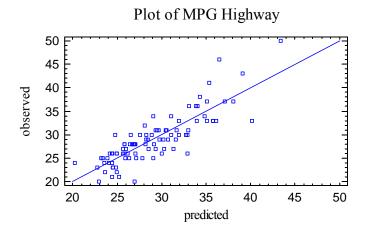
Further ANC	OVA for Variables in	the O	rder Fitted		_
Source	Sum of Squares	Df	Mean Square	F-Ratio	P-Value
Weight	1718.7	1	1718.7	207.47	0.0000
Weight^2	77.1615	1	77.1615	9.31	0.0030
Wheelbase	82.1713	1	82.1713	9.92	0.0022
Model	1878.03	3			

The table decomposes the model sum of squares SSR into contributions due to each coefficient by showing the increase in SSR as each term is added to the model. These sums of squares are often called *Type I sums of squares*. The F-Ratios compare the mean square for each term to the MSE of the fitted model. These sums of squares are useful when fitting polynomial models, as discussed in the *Polynomial Regression* documentation.

In the above table, all variables are statistically significant at the 1% significance level since their P-Values are well below 0.01.

Observed versus Predicted

The *Observed versus Predicted* plot shows the observed values of Y on the vertical axis and the predicted values \hat{Y} on the horizontal axis.



If the model fits well, the points should be randomly scattered around the diagonal line. Any change in variability from low values of Y to high values of Y might indicate the need to transform the dependent variable before fitting a model to the data. In the above plot, the variability appears to increase somewhat as the predicted values get large.

Residual Plots

As with all statistical models, it is good practice to examine the residuals. In a regression, the residuals are defined by

$$e_i = y_i - \hat{y}_i \tag{7}$$

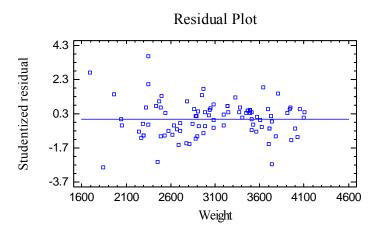
i.e., the residuals are the differences between the observed data values and the fitted model.

The Multiple Regression procedure creates 3 residual plots:

- 1. versus X.
- 2. versus predicted value \hat{Y} .
- 3. versus row number.

Residuals versus X

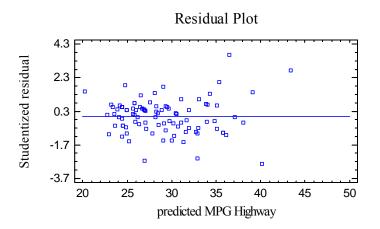
This plot is helpful in visualizing any missed curvature with respect to a selected variable.



No obvious curvature is detectable.

Residuals versus Predicted

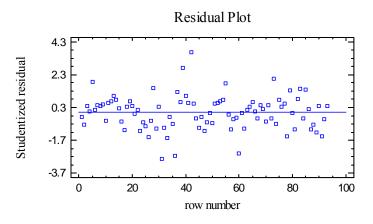
This plot is helpful in detecting any heteroscedasticity in the data.



Heteroscedasticity occurs when the variability of the data changes as the mean changes, and might necessitate transforming the data before fitting the regression model. It is usually evidenced by a funnel-shaped pattern in the residual plot. In the plot above, some increased variability in miles per gallon can be seen at high predicted values, which corresponds to the smaller cars. For the smaller cars, the miles per gallon appears to vary more than for the larger cars.

Residuals versus Observation

This plot shows the residuals versus row number in the datasheet:



If the data are arranged in chronological order, any pattern in the data might indicate an outside influence. In the above plot, no obvious trend is present, although there is a standardized residual in excess of 3.5, indicating that it is more than 3.5 standard deviations from the fitted curve.

Pane Options

Residual Plot Options	×
Plot Residuals Studentized Residuals Plot versus:	
Weight ² Weight ² Horsepower Wheelbase Drive Train = "front"	
OK Cancel Help	

- **Plot**: The following residuals may be plotted on each residual plot:
 - 1. Residuals the residuals from the least squares fit.
 - 2. Studentized residuals the difference between the observed values y_i and the predicted values \hat{y}_i when the model is fit using all observations except the *i*-th, divided by the estimated standard error. These residuals are sometimes called *externally deleted* residuals, since they measure how far each value is from the fitted model when that model is fit using all of the data except the point being considered. This is important, since a large outlier might otherwise affect the model so much that it would not appear to be unusually far away from the line.
- Plot versus: the independent variable to plot on the horizontal axis, if relevant.
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Unusual Residuals

Once the model has been fit, it is useful to study the residuals to determine whether any outliers exist that should be removed from the data. The *Unusual Residuals* pane lists all observations that have Studentized residuals of 2.0 or greater in absolute value.

Unusu	al Resi	duals		
		Predicted	Studentized	
Row	Y	Y	Residual	Residual
31	33.0	40.1526	-7.15265	-2.81
36	20.0	26.9631	-6.96309	-2.62
39	50.0	43.4269	6.5731	2.72
42	46.0	36.4604	9.53958	3.66
60	26.0	32.8753	-6.8753	-2.50
73	41.0	35.3266	5.67338	2.04

Studentized residuals greater than 3 in absolute value correspond to points more than 3 standard deviations from the fitted model, which is a very rare event for a normal distribution. In the sample data, row #42 is more 3.5 standard deviations out. Row #42 is a Honda Civic, which was listed in the dataset as achieving 46 miles per gallon, while the model predicts less than 37.

Points can be removed from the fit while examining any of the residual plots by clicking on a point and then pressing the *Exclude/Include* button on the analysis toolbar.

Influential Points

In fitting a regression model, all observations do not have an equal influence on the parameter estimates in the fitted model. Those with unusual values of the independent variables tend to have more influence than the others. The *Influential Points* pane displays any observations that have high influence on the fitted model:

		Mahalanobis	
Row	Leverage	Distance	DFITS
19	0.139122	13.5555	0.18502
28	0.246158	28.3994	0.685044
31	0.156066	15.6544	-1.225
36	0.0961585	8.58597	-0.849931
39	0.250016	29.0136	1.89821
50	0.0298891	1.78389	-0.463748
73	0.0352144	2.29596	0.451735
33	0.102406	9.27903	0.573505

Points are placed on this list for one of the following reasons:

- Leverage measures how distant an observation is from the mean of all *n* observations in the space of the *independent* variables. The higher the leverage, the greater the impact of the point on the fitted values \hat{y} . Points are placed on the list if their leverage is more than 3 times that of an average data point.
- **Mahalanobis Distance** measures the distance of a point from the center of the collection of points in the multivariate space of the independent variables. Since this distance is related to *leverage*, it is not used to select points for the table.

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• **DFITS** – measures the difference between the predicted values \hat{y}_i when the model is fit with and without the i-th data point. Points are placed on the list if the absolute value of DFITS exceeds $2p/\sqrt{n}$, where p is the number of coefficients in the fitted model.

In the sample data, rows #28 and #39 show a leverage value of nearly 6 times that of an average data point. Rows #31 and #39 have the largest values of DFITS. Removing high influence points is not recommended on a routine basis. However, it is important to be aware of their impact on the estimated model.

Confidence Intervals

The *Confidence Intervals* pane shows the potential estimation error associated with each coefficient in the model.

95.0% confidence intervals for coefficient estimates									
		Standard							
Parameter	Estimate	Error	Lower Limit	Upper Limit					
CONSTANT	55.0336	9.61121	35.9333	74.1339					
Weight	-0.023276	0.00475462	-0.0327248	-0.0138271					
Weight^2	0.00000230174	7.73643E-7	7.64284E-7	0.0000038392					
Wheelbase	0.220693	0.0860352	0.0497156	0.39167					

Pane Options

Confidence Intervals (ptions 🛛 🔀
	ОК
Confidence Level:	Cancel
, ^	Help

• Confidence Level: percentage level for the confidence intervals.

Correlation Matrix

The Correlation Matrix displays estimates of the correlation between the estimated coefficients.

Correlation ma	orrelation matrix for coefficient estimates								
	CONSTANT	Weight	Weight^2	Wheelbase					
CONSTANT	1.0000	-0.6247	0.7456	-0.6847					
Weight	-0.6247	1.0000	-0.9776	-0.1349					
Weight^2	0.7456	-0.9776	1.0000	-0.0508					
Wheelbase	-0.6847	-0.1349	-0.0508	1.0000					

This table can be helpful in determining how well the effects of different independent variables have been separated from each other. Note the high correlation between the coefficients of Weight and $Weight^2$. This is normal whenever fitting a non-centered polynomial and simply means that the coefficients could change dramatically if a different order polynomial was

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selected. The fact that the correlation between the coefficients of *Weight* and *Wheelbase* is small is more interesting, since it implies that there is little confounding between the estimated effects of those variables. Confounding or intermixing of the effects of two variables is a common problem when attempting to interpret models estimated from data that was not collected from a designed experiment.

Reports

The *Reports* pane creates predictions using the fitted least squares model. By default, the table includes a line for each row in the datasheet that has complete information on the X variables and a missing value for the Y variable. This allows you to add rows to the bottom of the datasheet corresponding to levels at which you want predictions without affecting the fitted model.

For example, suppose a prediction is desired for a car with a *Weight* of 3500 and a *Wheelbase* of 105. In row #94 of the datasheet, these values would be added but the *MPG Highway* column would be left blank. The resulting table is shown below:

Regression Results for MPG Highway									
	Fitted	Stnd. Error	Lower 95.0%	Upper 95.0%	Lower 95.0%	Upper 95.0%			
Row	Value	CL for Forecast	CL for Forecast	CL for Forecast	CL for Mean	CL for Mean			
94	24.7357	2.91778	18.9381	30.5333	23.7842	25.6872			

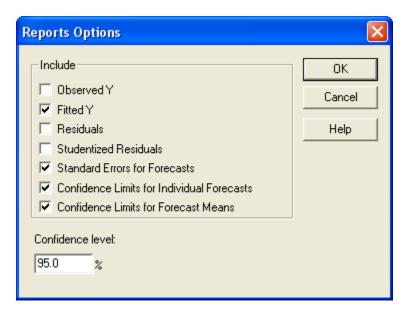
Included in the table are:

- **Row** the row number in the datasheet.
- Fitted Value the predicted value of the dependent variable using the fitted model.
- **Standard Error for Forecast** the estimated standard error for predicting a single new observation.
- **Confidence Limits for Forecast** prediction limits for new observations at the selected level of confidence.
- **Confidence Limits for Mean** confidence limits for the mean value of Y at the selected level of confidence.

For row #94, the predicted miles per gallon is 24.7. Models with those features can be expected to achieve between 18.9 and 30.5 miles per gallon in highway driving.

Using *Pane Options*, additional information about the predicted values and residuals for the data used to fit the model can also be included in the table.

Pane Options

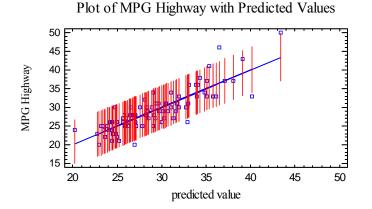


You may include:

- *Observed Y* the observed values of the dependent variable.
- *Fitted Y* the predicted values from the fitted model.
- *Residuals* the ordinary residuals (observed minus predicted).
- *Studentized Residuals* the Studentized deleted residuals as described earlier.
- *Standard Errors for Forecasts* the standard errors for new observations at values of the independent variables corresponding to each row of the datasheet.
- Confidence Limits for Individual Forecasts confidence intervals for new observations.
- *Confidence Limits for Forecast Means* confidence intervals for the mean value of Y at values of the independent variables corresponding to each row of the datasheet.

Interval Plots

The *Intervals Plots* pane can create a number of interesting types of plots. The plot below shows how precisely the miles per gallon of an automobile can be predicted.



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An interval is drawn on the plot for each observation in the dataset, showing the 95% prediction limits for a new observation at the corresponding predicted value.

Pane Options

Interval Plots Options	
Plot Limits For No Limits Predicted Values Means Forecasts Forecast Means Plot versus:	Confidence Level:
predicted values row number Weight Weight^2 Horsepower Wheelbase Drive Train = "front"	
OK Cancel	Help

- **Plot Limits For:** type of limits to be included. *Predicted Values* plots prediction limits at settings of the independent variables corresponding to each of the *n* observations used to fit the model. *Means* plots confidence limits for the mean value of Y corresponding to each of the n observations. *Forecasts* plots prediction limits for rows of the datasheet that have missing values for Y. *Forecast Means* plots confidence limits for the mean value of Y corresponding to each of Y corresponding to each row in the datasheet with a missing value for Y.
- **Plot Versus:** the value to plot on the horizontal axis.
- Confidence Level: the confidence percentage used for the intervals.

Autocorrelated Data

When regression models are used to fit data that is recorded over time, the deviations from the fitted model are often not independent. This can lead to inefficient estimates of the underlying regression model coefficients and P-values that overstate the statistical significance of the fitted model.

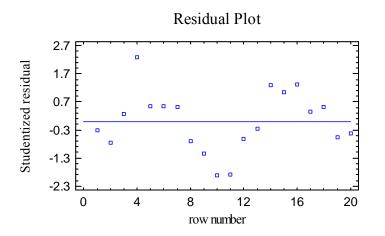
As an illustration, consider the following data from Neter et al. (1996), contained in the file *company.sgd*:

Year and quarter	Company Sales	Industry Sales
	(\$ millions)	(\$ millions)
1983: Q1	20.96	127.3
1983: Q2	21.40	130.0
1983: Q3	21.96	132.7
1983: Q4	21.52	129.4
1984: Q1	22.39	135.0
1984: Q2	22.76	137.1
1984: Q3	23.48	141.2
1984: Q4	23.66	142.8
1985: Q1	24.10	145.5
1985: Q2	24.01	145.3
1985: Q3	24.54	148.3
1985: Q4	24.30	146.4
1986: Q1	25.00	150.2
1986: Q2	25.64	153.1
1986: Q3	26.36	157.3
1986: Q4	26.98	160.7
1987: Q1	27.52	164.2
1987: Q2	27.78	165.6
1987: Q3	28.24	168.7
1987: Q4	28.78	171.7

Regressing *company sales* against *industry sales* resulting in a very good linear fit, with a very high R-squared:

Multiple Re	egression - c	company	v sale	S		
	able: company s					
Independent var	1 5					
industry sale						
	1	Standard	ł	Т		
Parameter	Estimate	Error		Statistic	P-Value	
CONSTANT	-1.45475	0.214146	6	-6.79326	0.0000	
industry sales	0.176283	0.001444	174	122.017	0.0000	
Analysis of Va	riance					
Source	Sum of Square	s Df	Mea	n Square	F-Ratio	P-Value
Model	110.257	1	110.	257	14888.14	0.0000
Residual	0.133302	18	0.00	740568		
Total (Corr.)	110.39	19				
Standard Error Mean absolute	sted for d.f.) = 9 of Est. = 0.0860 error = 0.069118 statistic = 0.734	563 36 4726 (P= 0 .0	0002)			

However, the Durbin-Watson statistic is very significant, and the estimated lag 1 residual autocorrelation equals 0.626. A plot of the residuals versus row number shows marked swings around zero:



Clearly, the residuals are not randomly distributed around the regression line.

To account for the autocorrelations of the deviations from the regression line, a more complicated error structure can be assumed. A logical extension of the random error model is to let the errors have a first-order autoregressive structure, in which the deviation at time t is dependent upon the deviation at time t-1 in the following manner:

$$y_t = \beta_0 + \beta_1 x_t + \varepsilon_t \tag{8}$$

$$\varepsilon_t = \rho \varepsilon_{t-1} + u_t \tag{9}$$

where $|\rho| < 1$ and u_t are independent samples from a normal distribution with mean 0 and standard deviation σ . In such a case, transforming both the dependent variable and independent variable according to

$$y'_{t} = y_{t} - \rho y_{t-1} \tag{10}$$

$$x_{t}' = x_{t} - \rho x_{t-1} \tag{11}$$

leads to the model

$$y'_{t} = \beta_{0} (1 - \rho) + \beta_{1} x'_{t} + u_{t}$$
(12)

which is a linear regression with random error terms.

The *Analysis Options* dialog box allows you to fit a model of the above form using the Cochrane-Orcutt procedure:

Multiple Regression Optio	ons				
Fitting Procedure C Ordinary Least Squares Forward Stepwise Selection Backward Stepwise Select Box-Cox Optimization Cochrane-Orcutt Optimization	Transform Power: 1.0 Addend: 0 Autocorr 0.626		, ,	OK Cancel Help	
I✓ Constant in Model	Constant in Model Selection C F-Ratio P-Value Max. Steps: 50			4.0	emove:

You may either specify the value of ρ in the *Autocorrelation* field and select *Ordinary Least Squares*, or select *Cochrane-Orcutt Optimization* and let the value of ρ will be determined iteratively using the specified value as a starting point. In the latter case, the following procedure is used:

Step 1: The model is fit using transformed values of the variables based on the initial value of ρ .

Step 2: The value of ρ is re-estimated using the values of ε_t obtained from the fit in Step 1.

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Step 3: Steps 1 and 2 are repeated between 4 and 25 times until the change in the derived value of ρ compared to the previous step is less than 0.01.

The results are summarized below using the sample data:

MULTER	•							
	<u>egression - c</u>		ny s	sale	<u>S</u>			
	able: company s	ales						
Independent va								
industry sale	es							
~ . ~	a i							
Cochrane-Orcu	tt transformation			tocor		765941		
		Standa	ırd		Т			
Parameter	Estimate	Error			Statistic	P-Value	?	
CONSTANT	-0.64188	0.6426	0.642688		-0.998744	0.3319		
industry sales	0.171085	0.0040	895	8	41.8343	0.0000	0.0000	
Analysis of Va						i		
Source	Sum of Square	s D	f 1	Mean	n Square	F-Ratio	P-Value	
Model	7.7425	1	7	7.742	25	1750.11	0.0000	
Residual	0.0752083	17	' (0.004	42402			
Total (Corr.)	7.81771	18						
· · · · · ·	•							
R-squared $=$ 99	.038 percent							
R-squared (adju	usted for $d.f.$ = 9	98.9814	perco	ent				
Standard Error	of Est. $= 0.0665$	133						
Mean absolute	error = 0.053173	31						
Durbin-Watson	statistic $= 1.735$	54						
Lag 1 residual a	autocorrelation =	= 0.0990	69					
-								

The above output shows that, at the final value of $\rho = 0.766$, the Durbin-Watson statistic and the lag 1 residual autocorrelation, computed using the residuals from the regression involving the transformed variables, are much more in line with that expected if the errors are random. The model also changed somewhat.

Save Results

The following results may be saved to the datasheet:

- 1. *Predicted Values* the predicted value of Y corresponding to each of the *n* observations.
- 2. Standard Errors of Predictions the standard errors for the *n* predicted values.
- 3. Lower Limits for Predictions the lower prediction limits for each predicted value.
- 4. Upper Limits for Predictions the upper prediction limits for each predicted value.
- 5. *Standard Errors of Means* the standard errors for the mean value of Y at each of the *n* values of X.
- 6. *Lower Limits for Forecast Means* the lower confidence limits for the mean value of Y at each of the *n* values of X.
- 7. *Upper Limits for Forecast Means* the upper confidence limits for the mean value of Y at each of the *n* values of X.
- 8. Residuals the *n* residuals.
- 9. *Studentized Residuals* the *n* Studentized residuals.
- 10. Leverages the leverage values corresponding to the *n* values of X.
- 11. DFITS Statistics the value of the DFITS statistic corresponding to the *n* values of X.
- 12. *Mahalanobis Distances* the Mahalanobis distance corresponding to the *n* values of X.

Calculations

Regression Model

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k$$
(13)

Error Sum of Squares

Unweighted:
$$SSE = \sum_{i=1}^{n} \left(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i - \hat{\beta}_2 x_2 - \dots - \hat{\beta}_k x_k \right)^2$$
 (14)

Weighted:
$$SSE = \sum_{i=1}^{n} w_i \left(y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i - \hat{\beta}_2 x_2 - \dots - \hat{\beta}_k x_k \right)^2$$
 (15)

Coefficient Estimates

 $\hat{\beta} = (X W X)^{-1} (X W Y) \tag{16}$

$$s^{2}\left\{\hat{\boldsymbol{\beta}}\right\} = MSE(X'WX)^{-1}$$
(17)

$$MSE = \frac{SSE}{n-p} \tag{18}$$

where $\hat{\beta}$ is a column vector containing the estimated regression coefficients, X is an (n, p) matrix containing a 1 in the first column (if the model contains a constant term) and the settings of the k predictor variables in the other columns, Y is a column vector with the values of the dependent variable, and W is an (n, n) diagonal matrix containing the weights w_i on the diagonal for a weighted regression or 1's on the diagonal if weights are not specified. A modified sweep algorithm is used to solve the equations after centering and rescaling of the independent variables.

Analysis of Variance

With constant term:

Source	Sum of Squares	Df	Mean	F-Ratio
Model	$SSR = b'X'WY - \frac{\left(\sum_{i=1}^{n} w_i y_i\right)^2}{\sum_{i=1}^{n} w_i}$	k	$Square$ $MSR = \frac{SSR}{k}$	$F = \frac{MSR}{MSE}$
Residual	SSE = Y'WY - b'X'WY	n-k-1	$MSE = \frac{SSE}{n-k-1}$	
Total (corr.)	$SSTO = \sum_{i=1}^{n} w_i (y_i - \overline{y})^2$	n-1		

Without constant term:

Source	Sum of Squares	Df	Mean Square	F-Ratio
Model	SSR = b'X'WY	k	$MSR = \frac{SSR}{k}$	$F = \frac{MSR}{MSE}$
Residual	SSE = Y'WY - b'X'WY	n-k	$MSE = \frac{SSE}{n-k}$	
Total	SSTO = Y'WY	n		

R-Squared

$$R^{2} = 100 \left(\frac{SSR}{SSR + SSE} \right) \%$$
⁽¹⁹⁾

Adjusted R-Squared

$$R_{adj}^{2} = 100 \left[1 - \left(\frac{n-1}{n-p} \right) \frac{SSE}{SSR + SSE} \right] \%$$
(20)

Stnd.Error of Est.

$$\hat{\sigma} = \sqrt{MSE} \tag{21}$$

Residuals

$$e_{i} = y_{i} - \hat{\beta}_{o} - \hat{\beta}_{1}x_{1} - \dots - \hat{\beta}_{k}x_{k}$$
(22)

Mean Absolute Error

...

$$MAE = \frac{\sum_{i=1}^{n} w_i |e_i|}{\sum_{i=1}^{n} w_i}$$
(23)

Durbin-Watson Statistic

$$D = \frac{\sum_{i=2}^{n} (e_i - e_{i-1})^2}{\sum_{i=1}^{n} e_i^2}$$
(24)

If n > 500, then

$$D^* = \frac{|D-2|}{\sqrt{4/n}}$$
(25)

is compared to a standard normal distribution. For $100 < n \le 500$, D/4 is compared to a beta distribution with parameters

$$\alpha = \beta = \frac{n-1}{2} \tag{26}$$

For smaller sample sizes, D/4 is compared to a beta distribution with parameters which are based on the trace of certain matrices related to the *X* matrix, as described by Durbin and Watson (1951) in section 4 of their classic paper.

Lag 1 Residual Autocorrelation

$$r_{1} = \frac{\sum_{i=2}^{n} e_{i} e_{i-1}}{\sum_{i=1}^{n} e_{i}^{2}}$$
(27)

Leverage

$$h_i = diag \left\{ X_i' (X'WX)^{-1} X_i \right\} w_i$$
(28)

$$\overline{h} = \frac{p}{n} \tag{29}$$

Studentized Residuals

$$d_i = \frac{e_i \sqrt{w_i}}{\sqrt{MSE_i (1 - h_i)}}$$
(30)

Mahalanobis Distance

$$MD_{i} = \left(\frac{h_{i} - w_{i} / \sum_{i=1}^{n} w_{i}}{1 - h_{i}}\right) \frac{n(n-2)}{n-1}$$
(31)

DFITS

$$DFITS_{i} = \frac{d_{i}}{\sqrt{w_{i}}} \sqrt{\left(\frac{h_{i}}{1-h_{i}}\right)}$$
(32)

Standard Error for Forecast

$$s\{Y_{h(new)}\} = \sqrt{MSE(1 + X_{h}'(X'WX)^{-1}X_{h})}$$
(33)

Confidence Limit for Forecast

 $\hat{Y}_h \pm t_{\alpha/2, n-p} s \left\{ Y_{h(new)} \right\}$ (34)

Confidence Limit for Mean

$$\hat{Y}_{h} \pm t_{\alpha/2,n-p} \sqrt{MSE\left(X_{h}'(XWX)^{-1}X_{h}\right)}$$
(35)